Math 6070-1: Spring 2013 Problem set 5

Due date: April 24, 2013

- 1. (Normal-data bootstrapping) Simulate 1,000 i.i.d. samples from N(0,1), and let us call this sample X_1, \ldots, X_{1000} .
 - (a) Take N bootstrap subsamples in order to estimate

 $\theta = \mathcal{P}\left\{\cosh(\bar{X}) \le 3\right\},\,$

for N = 1,000, 5,000, and 10,000.

- (b) Use 50,000 simulations [each of size 1,000, of couse!] in order to find an accurate evaluation of θ .
- (c) Discuss how well your bootstrap method is working.
- 2. (Introduction to time series) Download the atmospheric CO_2 concentration data by Uhse, Schmidt, and Levin from

http://cdiac.ornl.gov/ftp/trends/co2/westerland.co2.

- (a) This data is a time series of the form (t, x_t) , where t = 1, 2, ... Plot the time series [that is, plot t versus x_t for all times t].
- (b) Define a new data set, starting with $y_1 := x_1$, and then

 $y_t := x_t - x_{t-1} \qquad \text{for all } t \ge 2.$

Plot t versus y_t . Justify the phrase, "we have obtained the y_t 's by de-trending the x_t 's."

(c) In order to understand more deeply what you have done, posit a statistical model of the type

$$x_t = \beta_0 + \beta_1 t + w_t \quad \text{for all } t \ge 1,$$

where w_t is an i.i.d. sequence of random variables with mean zero and unknown variance σ^2 ; this is called a *linear regression model*. The constants β_0 and β_1 are unknown statistical parameters.

i. Compute y_t explicitly in terms of w_t .

- ii. Are the w_t 's independent? Perform a test, and describe both your methods, and the outcome.
- iii. Describe an unbiased estimator for β_0 , using only the y_t 's. Prove that your estimator is unbiased. You may not use, without proof, any facts from regression analysis.
- (d) Let X_1, \ldots, X_n be i.i.d. $N(\theta, 1)$, where θ is the unknown parameter. Define T_n to be the *Hodges estimator* of θ ; that is,

$$T_n := \begin{cases} \frac{1}{2}\bar{X} & \text{if } |\bar{X}| \le n^{-1/4}, \\ \bar{X} & \text{if } |\bar{X}| > n^{-1/4}. \end{cases}$$

- i. Is T_n biased? Analyze carefully.
- ii. Prove that T_n is consistent for θ by proving the even more general statement that

$$\sqrt{n} (T_n - \theta) \stackrel{d}{\longrightarrow} \mathcal{N}(0, \sigma^2(\theta)) \quad \text{as } n \to \infty,$$

where

$$\sigma^{2}(\theta) := \begin{cases} 1/4 & \text{if } \theta = 0, \\ 1 & \text{if } \theta \neq 0. \end{cases}$$

iii. Let F_n^{θ} denote the cdf of $\sqrt{n}(T_n - \theta)$. That is,

$$F_n^{\theta}(x) := \mathbf{P}\left\{\sqrt{n}\left(T_n - \theta\right) \le x\right\}.$$

The previous exercise can be restated as follows: $F_n^{\theta}(x) \to \Phi(x/\sigma(\theta))$ for all $x \in \mathbf{R}$ as $n \to \infty$, where Φ denotes the standard normal cdf. Prove the stronger assertion that

$$\max_{x \in \mathbf{R}} |F_n^{\theta}(x) - \Phi(x/\sigma(\theta))| \to 0 \quad \text{as } n \to \infty.$$

iv. (Hard/extra credit) The parametric bootstrap estimator of $F_n^\theta(x)$ is, of course, $F_n^{\bar{X}}(x).$ Prove that

$$\max_{x \in \mathbf{R}} \left| F_n^{\bar{X}}(x) - F_n^{\theta}(x) \right| \stackrel{\mathrm{P}}{\longrightarrow} 0 \qquad \text{as } n \to \infty,$$

if and only if $\theta = 0$. In other words, the bootstrap fails generically in this case, unless θ happens to be zero.