Probability Prelim Exam

August 2013

Read the following instructions before you begin:

- You may attempt all of 10 problems in this exam. However, you can turn in solutions for **at most** 6 problems. On the outside of your exam booklet, indicate which problem you are turning in.
- Each problem is worth 10 points; 40 points or higher will result in a pass.
- If you think that a problem is misstated, then interpret that problem in such a way as to make it nontrivial.
- If you wish to quote a result that was not in your 6040 text, then you need to carefully state and prove that result.

Exam problems begin here:

Throughout, (Ω, \mathcal{F}, P) denotes the underlying probability space.

- 1. Let X, X_1, X_2, \ldots be independent random variables, all distributed according to the standard normal distribution.
 - (a) Prove that there exist finite constants $c_1, c_2, x_0 > 0$, such that

$$\frac{c_1}{x} \mathrm{e}^{-x^2/2} \leqslant \mathrm{P}\{X > x\} \leqslant \frac{c_2}{x} \mathrm{e}^{-x^2/2} \qquad \text{for all } x \geqslant x_0.$$

(b) Use the inequalities of part (a) in order to prove that if X_1, X_2, \ldots are independent standard normal random variables then

$$\limsup_{n \to \infty} \frac{X_n}{\sqrt{2 \ln n}} = -\liminf_{n \to \infty} \frac{X_n}{\sqrt{2 \ln n}} = 1 \qquad \text{a.s.}$$

2. Let X_1, X_2, \ldots be i.i.d. with $P\{X_1 = 1\} = \frac{1}{4}$ and $P\{X_1 = 0\} = \frac{3}{4}$. Let N denote the smallest integer $k \ge 2$ such that $X_{k-1} = 0, = X_k = 1$; that is, N is the first time the pattern "01" occurs in the infinite sequence X_1, X_2, \ldots . Compute E(N). 3. Suppose X_1, X_2, \ldots, X_n are i.i.d. random variables, and $1 \leq X_1 \leq 2$ a.s. Compute, for every $1 \leq i \leq j \leq n$, the quantity

$$\operatorname{E}\left(\frac{X_i+\cdots+X_j}{X_1+\cdots+X_n}\right).$$

4. Suppose μ is a probability measure on the Borel subsets of (0, 1) such that $\int |\ln(x)| \, \mu(dx) < \infty$. Prove that

$$I(n) := \left(\int x^{1/n} \, \mu(\mathrm{d}x) \right)^n \to \exp\left(\int \ln(x) \, \mu(\mathrm{d}x) \right) \qquad \text{as } n \to \infty.$$

[Hint: $I(n) = \mathbb{E}(\prod_{j=1}^{n} X_j^{1/n})$, where X_1, \ldots, X_n are i.i.d.]

- 5. Let $\mathbf{X}_n := (X_{1,n}, \dots, X_{d,n})$ be a random variable in \mathbf{R}^d for every $n \ge 1$. Let $\mathbf{Y} := (Y_1, \dots, Y_d)$ be another random variable in \mathbf{R}^d . Prove that $\mathbf{X}_n \Rightarrow \mathbf{Y}$ if and only if the 1-dimensional random variable $\sum_{i=1}^d \lambda_i X_{i,n}$ converges weakly to $\sum_{i=1}^d \lambda_i Y_i$ for every $\mathbf{\lambda} := (\lambda_1, \dots, \lambda_d) \in \mathbf{R}^d$.
- 6. For every random vector (X, Y) in \mathbf{R}^2 we can define a function \mathscr{Q} of two variables as follows:

$$\mathscr{Q}(a,b) := \mathbf{P}\{X > a, Y > b\} - \mathbf{P}\{X > a\}\mathbf{P}\{Y > b\}.$$

(a) Suppose $f, g : \mathbf{R} \to \mathbf{R}$ are bounded and have integrable, bounded, and continuous derivatives. Then prove that

$$\mathbf{E}[f(X)g(Y)] - \mathbf{E}[f(X)]\mathbf{E}[g(Y)] = \iint_{\mathbf{R}^2} \mathscr{Q}(a,b)f'(a)g'(b)\,\mathrm{d}a\,\mathrm{d}b.$$

(b) Suppose $\mathscr{Q}(a,b) \ge 0$ for all $a, b \in \mathbf{R}$. Prove that if X and Y have two finite moments, then $\operatorname{Cov}(X,Y) \ge 0$ and there exists a finite constant α , depending only on f and g, such that

$$|\mathbf{E}[f(X)g(Y)] - \mathbf{E}[f(X)]\mathbf{E}[g(Y)]| \leq \alpha \mathrm{Cov}(X, Y).$$

[In other words, if $\mathcal{Q} \ge 0$ and X and Y have a small covariance, then X and Y are almost independent.]

7. Suppose that, for every integer $n \ge 1$, X(n) has a Gamma distribution with parameters n + 1 and 1; that is, the probability density function of X(n) is

$$f_n(x) := \frac{x^n e^{-x}}{n!} \mathbf{1}_{(0,\infty)}(x).$$

(a) Compute E(X(n)) and Var(X(n)) for every integer $n \ge 1$.

(b) Prove that, as $n \to \infty$,

$$\frac{X(n) - n}{\sqrt{n}} \Rightarrow \mathcal{N}(0, 1).$$

8. Let $\{W(t)\}_{t\geq 0}$ be a standard Brownian motion, and define

$$U(t) := e^{-t/2} W(e^t) \quad \text{for all } t \ge 0.$$

- (a) Prove that $\{U(t)\}_{t\geq 0}$ is Gaussian process; compute its mean and covariance functions.
- (b) Prove that

$$\limsup_{t \to \infty} \frac{U(t)}{\sqrt{2 \ln t}} = -\liminf_{t \to \infty} \frac{U(t)}{\sqrt{2 \ln t}} = 1 \qquad \text{a.s.}$$

9. Let X be an integer-valued random variable that is non-degenerate; this means that there exist at least two integers n, m such that $P\{X = n\}$ and $P\{X = m\}$ are both strictly positive. Prove that the set

$$\mathcal{V} := \left\{ t \in \mathbf{R} : |\mathbf{E}[\mathbf{e}^{itX}]| = 1 \right\}$$

has zero Lebesgue measure. [Hint: If $t \in \mathcal{V}$ then $Z := e^{it(X-Y)}$ has mean one, where Y is an independent copy of X. Proceed by proving that $E(|Z-1|^2) = 0.$]

10. Consider i.i.d. X_1, X_2, \ldots that are strictly positive, and define

$$S_n := \sqrt{X_1 + \sqrt{X_2 + \sqrt{X_3 + \sqrt{\dots + \sqrt{X_n}}}}}, \quad \text{for all } n \ge 1.$$

In other words, if $f(a, b) := \sqrt{a + \sqrt{b}}$ for all a, b > 0, then $S_1 := f(X_1, 0)$ and $S_{k+1} := f(S_k, X_{k+1})$ for all $k \ge 0$.

Prove that $\lim_{n\to\infty} S_n$ exists a.s. and is finite a.s.