Name

Suppose 100 students from the University of Utah will participate in an NZT trial. Each student will take an IQ test without NZT and then again on a different day with NZT. For simplicity, suppose we know that the true standard deviation of differences is 10. Using a two-sided alternative to test the null hypothesis that NZT does not improve IQ, find the power of the test if NZT actually improves IQ by 1 point (on average). Note that this is all calculated prior to conducting the experiment.

$$D_i \wedge iid (M, 10^2)$$

Reject if $\frac{\overline{d}}{10} = \overline{d}$ is not between $\frac{z_{H_2}}{z_{H_2}}$ and $\frac{z_{I-\frac{w}{2}}}{\sqrt{100}}$

$$\pi(1) = P(\overline{D} < z_{w_{2}} | m=1) + P(\overline{D} > \overline{z_{1-w_{2}}} | m=1)$$

$$= P(\overline{D} - 1 < z_{w_{2}} - 1 | m=1)$$

$$+ P(\overline{D} - 1 > \overline{z_{1-w_{2}}} - 1 | m=1)$$

$$\approx \Phi(\overline{z_{w_{2}}} - 1) + 1 - \overline{\Phi}(\overline{z_{1-w_{2}}} - 1)$$

(11.19 in the textbook)

19. Consider independent random samples from two normal distributions, $X_i \sim N(\mu_1, \sigma_1^2)$ and $Y_j \sim N(\mu_2, \sigma_2^2)$; $i = 1, ..., n_1, j = 1, ..., n_2$. Assuming that μ_1 and μ_2 are known, derive a $100(1 - \alpha)\%$ confidence interval for σ_2^2/σ_1^2 based on sufficient statistics.

$$\frac{(n_{i}-1)S_{i}^{z}}{\sigma_{i}^{z}} \sim \mathcal{F}^{z}(n_{i}-1)$$

$$\frac{(n_{2}-1)S_{z}^{2}}{\sigma_{z}^{2}} \sim \gamma^{2}(n_{z}-1)$$

$$\frac{(n_{1}-1)S_{1}^{2}}{(n_{2}-1)S_{2}^{2}} \cdot \frac{\sigma_{2}^{2}}{\sigma_{1}^{2}} \cdot \frac{(n_{2}-1)}{(n_{1}-1)} = \frac{S_{1}^{2}\sigma_{2}^{2}}{S_{2}^{2}\sigma_{1}^{2}} \sim F(n_{1}-1, n_{2}-1)$$

$$-\alpha = P\left(f(n_{1}-1, n_{2}-1) < \frac{S_{1}^{2}\sigma_{2}^{2}}{S_{2}^{2}\sigma_{1}^{2}} < f_{1-\sigma_{2}^{2}} (n_{1}-1, n_{2}-1)\right)$$

$$= P\left(\frac{S_{2}^{2}}{S_{1}^{2}} + \frac{S_{2}^{2}\sigma_{1}^{2}}{\sigma_{1}^{2}} + \frac{\sigma_{2}^{2}}{\sigma_{1}^{2}} < \frac{\sigma_{2}^{2}}{\sigma_{1}^{2}} < \frac{S_{2}^{2}}{\sigma_{1}^{2}} + \frac{S_{2}^{2}\sigma_{1}^{2}}{S_{1}^{2}} + \frac{S_{2}^{2}\sigma_{1}^{2}}{\sigma_{1}^{2}} < \frac{S_{2}^{2}}{\sigma_{1}^{2}} + \frac{S_{2}^{2}\sigma_{1}^{2}}{S_{1}^{2}} + \frac{S_{2}^{2}\sigma_{1}^{2}}{S_$$

A
$$100(1-\alpha)$$
% CI for $\frac{\sigma_{z}^{2}}{\sigma_{1}^{2}}$ is:

$$\int \left(\frac{s_{z}^{2}}{s_{1}^{*}} \int_{\sigma_{z}} (n_{1}-1, n_{z}-1) \int_{\sigma_{z}} \frac{s_{z}^{2}}{s_{1}^{*}} \int_{1-\alpha_{z}} (n_{1}-1, n_{z}-1) \right)$$

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19. Let X₁,..., X_n be a random sample from a normal distribution, X₁ ~ N(μ, 1).
(a) Find a UMP test of H₀: μ = μ₀ against H_a: μ < μ₀.

Suppose you have an EXP(θ , η) population. Derive a GLR test of the null hypothesis that θ = 4 against the alternative that θ > 4. Write down the test statistic and all associated maximum likelihood estimates. Then state the condition under which you will reject the null hypothesis in order to have a test of approximate size 0.05.

$$\hat{\Theta} = \max(\bar{x} - x_{1:n}, 4)$$

$$\hat{\chi} = x_{1:n}$$

$$\hat{\Theta}_{0} = 4$$

$$\hat{\chi}_{0} = x_{1:n}$$

$$-\frac{(x_{1} - x_{1:n})}{4}$$

$$\hat{\chi} = \prod_{i=1}^{n} \frac{1}{4} e$$

$$\frac{\prod_{i=1}^{n} \frac{1}{4} e}{\prod_{i=1}^{n} (\max(\bar{x} - x_{1:n}, 4))} e^{-\frac{(x_{i} - x_{1:n})}{\max(\bar{x} - x_{1:n}, 4)}}$$

Reject Ho if $-2\ln(\lambda) > \chi^2_{0.95}(1)$.

The following table shows the number of m&ms of each type and color that were eaten by 3rd graders in an experiment.

	Red	Blue	Green	Yellow
Peanut	13 (🖈)	17	11	16
Peanut butter	14	19	14	15
Almond	15	16	17	19
Plain	13	17	17	16

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Test whether there is a relationship between type and color of m&m. In particular, what is the expected count of red peanut m&ms, and what is the p-value? You don't need to write out the entire test statistic.

Expected count of red peanut mems is
$$\left(\frac{13+14+15+13}{n}\right)\left(\frac{13+17+11+16}{n}\right)n$$
, where n is the sym

of all counts in the table.
Expected count =
$$\frac{55}{249} \cdot \frac{57}{349} \cdot \frac{249}{349} = 12,59$$
.

$$P-value = P(\chi^{2}(9) \ge t) \quad \text{where}$$
$$t = \sum \frac{(o_{i} - e_{i})^{2}}{e_{i}}.$$

The following table shows the number of m&ms of each type and color that were eaten by 3rd graders in an experiment.

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Peanut butter	14	-19	14	15
Almond	15	16	17 Total loss and the second s	19
Plain	13	17	17	16

Test whether there the proportion of blue m&ms is the same for Plain and Peanut m&ms. In particular, what is the expected count of plain blue m&ms, and what is the p-value? You don't need to write out the entire test statistic.

	Blue	Non-Blue	
lean ut	17	13+11+16=40	and a second second second
Plain	17	13+17+16=46	-

$$\hat{p} = \frac{17+17}{\text{total}} = \frac{17+17}{17+17+40+46} = \frac{34}{130} = .2615$$

Expected plain blues =
$$(17+46)\frac{34}{130} = 16.48$$

$$p-value = P(\chi^2(1) \ge t)$$

where t is the outcome of the test statistic.

Circle only the statements that are true. You will get 1 point for 6 correct answers; 2 points for 7 correct; 3 points for 8 correct; 4 points for 9 correct. In other words, you will get max(# correct -5, 0) points for this problem.

- 1. The outcome of a random interval that contains the parameter of interest with 95% probability is called a confidence interval for the parameter of interest.
- 2.) If two researchers use the exact same data from an experiment and report different confidence intervals for the same parameter, then they may both be right.
- 3. A one-sided upper confidence interval is most appropriate when a drug company wants to establish (with 95% confidence) that no more than 5% of people will have an allergic reaction to a birth control patch that they are developing.
- (4.) MLE θ is a pivotal quantity for θ if MLE is a maximum likelihood estimator for θ and θ is a location parameter.
- 5. Pivotal quantities are functions of the random sample and the parameter of interest.
- 6. If (1, 2) is a 95% confidence interval for μ , then (1, 4) is a 95% confidence interval for μ^2 .
- 7. If the null hypothesis is not rejected, then you should conclude that the alternative hypothesis is plausible.
- 8. The Neyman-Pearson lemma provide a method to obtain the most powerful test of a simple null hypothesis against a two-sided alternative.
- 9. The generalized likelihood ratio test provides a reasonable test, whose test statistic is the likelihood ratio, and whose distribution under the null hypothesis is $\chi^2(r)$, where r parameters are fixed under the null.