

1. Let X_1, X_2 and X_3 be three independent random variables. X_1 is $\chi^2(3)$ and X_2 is $\chi^2(2)$ and X_3 is $\chi^2(2)$. Compute

$$P\{X_1/(X_2 + X_3) \leq 1\}$$

using one of the enclosed tables.

$$P\left(\frac{X_1}{X_2 + X_3} \leq 1\right) = P\left(\frac{X_1/3}{\frac{X_2 + X_3}{4}} \leq \frac{4}{3}\right)$$

$$= P\left(Y \leq 4/3\right) \text{ where } Y \sim F(3, 4)$$

2. Let X_1, X_2, \dots, X_n be independent identically distributed random variables with density function

$$h(t; \theta) = \begin{cases} 0, & \text{if } t \notin [0, 1] \\ \theta t^{\theta-1}, & \text{if } t \in [0, 1]. \end{cases}$$

Find a sufficient statistic for θ .

$$\begin{aligned} f(x; \theta) &= \prod \theta x_i^{\theta-1} \mathbb{1}_{\{x_i \in [0, 1]\}} \\ &= \theta^n e^{(\theta-1) \sum \log(x_i)} \mathbb{1}_{\{x_i \in [0, 1]\}}. \end{aligned}$$

Therefore $\sum \log(x_i)$ is sufficient for θ .

$\prod x_i$ is also sufficient for θ .

3. Let X_1, X_2, \dots, X_n be independent identically distributed random variables with density function

$$h(t; \theta) = \begin{cases} 0, & \text{if } t \notin [0, \theta] \\ \frac{2t}{\theta^2}, & \text{if } t \in [0, \theta]. \end{cases}$$

Find a sufficient statistic for θ .

$$\begin{aligned} f(\underline{x}; \theta) &= \prod \frac{2x_i}{\theta^2} \mathbb{1}\{x_i \in [0, \theta]\} \\ &= \frac{2^n}{\theta^{2n}} \left(\prod x_i \right) \mathbb{1}\{x_{(n)} \leq \theta\} \mathbb{1}\{x_{(1)} \geq 0\}. \end{aligned}$$

Thus $X_{(n)}$ is sufficient for θ .

4. Let X_1, X_2, \dots, X_n be independent identically distributed random variables with density function

$$h(t; \theta) = \frac{1}{2\theta} e^{-|t|/\theta}, \quad -\infty < t < \infty.$$

Find the maximum likelihood estimator for θ and compute the asymptotic variance of the maximum likelihood estimator.

$$L(\theta) = \frac{1}{2^n \theta^n} e^{-\frac{1}{\theta} \sum |x_i|}, \quad \text{which is maximized}$$

$$\text{at } \hat{\theta} = \frac{\sum |x_i|}{n}.$$

The asymptotic variance of the unique MLE is the CRLB, $\frac{\theta^2}{n}$.

5. Let X_1, X_2, \dots, X_n be independent identically distributed random variables with density function

$$h(t; \theta) = \begin{cases} 0, & \text{if } t \notin [0, \theta] \\ \frac{2t}{\theta^2}, & \text{if } t \in [0, \theta]. \end{cases}$$

Find the maximum likelihood estimator for θ and compute the bias of the maximum likelihood estimator for θ .

$$\begin{aligned} L(\theta) &= \frac{2^n \theta^{-2n}}{\cdot} \prod x_i \mathbb{1}\{x_i \in [0, \theta]\} \\ &= 2^n \theta^{-2n} (\prod x_i) \mathbb{1}\{x_{n:n} \leq \theta\}, \end{aligned}$$

which is maximized at $\hat{\theta} = x_{n:n}$.

$$\text{Bias}(\hat{\theta}) = E(\hat{\theta}) - \theta = \frac{-\theta}{2n+1}$$

6. Let X_1, X_2, \dots, X_n be independent identically distributed random variables with density function

$$h(t; \theta) = \begin{cases} 0, & \text{if } -\infty < t < 1 \\ \theta t^{-(\theta+1)}, & \text{if } 1 \leq t < \infty \end{cases}$$

Find a moment estimator for θ .

$$\bar{X} = \int_1^{\infty} x \theta x^{-(\theta+1)} dx = \frac{-\theta}{1-\theta} \quad \text{for } \theta > 1$$

$$\text{Thus } \hat{\theta} = \frac{\bar{X}}{\bar{X} - 1}$$

7. Let X_1, X_2 and X_3 be three independent normal random variables. X_1 is normal $N(0, 4)$, X_2 is normal $N(0, 12)$ and X_3 is normal $N(0, 9)$. Compute

$$P\left\{\frac{X_1 + X_2}{|X_3|} \leq 3\right\}$$

using one of the enclosed tables.

$$P\left(\frac{X_1 + X_2}{|X_3|} \leq 3\right) = P\left(\frac{\frac{X_1 + X_2}{4}}{\sqrt{\frac{X_3^2}{3^2}}} \leq \frac{3}{4} \cdot 3\right)$$

$$= P\left(Y \leq \frac{9}{4}\right) \quad \text{where} \quad Y \sim T(1).$$

8. Let X_1, X_2, \dots, X_n be independent identically distributed random variables with probability mass function

$$f(t; \theta) = \frac{2!}{t!(2-t)!} \theta^t (1-\theta)^{2-t}, \text{ if } t = 0, 1, 2.$$

Find the uniformly minimum variance estimator for θ .

$$\begin{aligned} f(\underline{x}; \theta) &= \prod \binom{2}{x_i} \theta^{x_i} (1-\theta)^{2-x_i} \mathbb{1}_{\{x_i \in \{0, 1, 2\}\}} \\ &= \left(\prod \binom{2}{x_i} \mathbb{1}_{\{x_i \in \{0, 1, 2\}\}} \right) \theta^{\sum x_i} (1-\theta)^{2n - \sum x_i}. \end{aligned}$$

Thus $\sum X_i$ is sufficient for θ .

$\sum X_i \sim \text{BIN}(2n, \theta)$, which is complete since it is a member of the REC.

Thus $\frac{\sum X_i}{2n}$ is a UMVUE for θ .