

Math 6070-1, Spring 2014; Partial Solutions to Assignment #2

1. Suppose X_1, \dots, X_n is a random [that is, i.i.d.] sample from a $\text{Uniform}(\theta, 2\theta)$ distribution, where $\theta > 0$ is an unknown parameter.

- (a) Find the maximum likelihood estimator for θ .

Solution. The likelihood function is

$$L(\theta) = \prod_{j=1}^n \theta^{-1} \mathbf{I}\{\theta < X_j < 2\theta\} = \theta^{-n} \mathbf{I}\left\{\frac{X_{n:n}}{2} \leq \theta \leq X_{1:n}\right\},$$

where $X_{j:n}$ denotes the j th order statistic of $\{X_1, \dots, X_n\}$. Because the function $h(\theta) := \theta^{-n}$ is decreasing, it follows that the MLE is

$$\hat{\theta} = \frac{X_{n:n}}{2} \quad \text{provided that} \quad \frac{X_{n:n}}{2} \leq X_{1:n}.$$

It remains to check that we always have $X_{n:n}/2 \leq X_{1:n}$, so that the MLE is always uniquely equal to $X_{n:n}/2$. [This follows simply because $X_{n:n} \leq 2\theta$ therefore, $X_{n:n}/2 \leq \theta$, whereas $X_{1:n} \geq \theta$.]

- (b) Prove that the MLE is consistent.

Solution. Recall that $\hat{\theta} = X_{n:n}/2$. If $x > 0$, then

$$\mathbf{P}\{\hat{\theta}/\theta \leq x\} = (\mathbf{P}\{X_1 \leq 2\theta x\})^n.$$

[If $x < 0$ then $\mathbf{P}\{\hat{\theta}/\theta \leq x\} = 0$, vacuously.] The last displayed probability is equal to one if $2\theta x \geq 2\theta$ [i.e., $x \geq 1$] and zero if $2\theta x \leq \theta$ [i.e., $x \leq 1/2$]. On the other hand, if $\frac{1}{2} < x < 1$, then it follows that

$$\mathbf{P}\{\hat{\theta}/\theta \leq x\} = \left(\int_{\theta}^{2\theta x} \theta^{-1} dy\right)^n = (2x - 1)^n.$$

In conclusion, the pdf of $\hat{\theta}/\theta$ is

$$f_{\hat{\theta}/\theta}(x) = \begin{cases} 2n(2x - 1)^{n-1} & \text{if } \frac{1}{2} < x < 1, \\ 0 & \text{otherwise.} \end{cases}$$

We can use this to compute

$$\begin{aligned}
E[\hat{\theta}/\theta] &= 2n \int_{1/2}^1 x(2x-1)^{n-1} dx \\
&= n \int_0^1 \left(\frac{y+1}{2}\right) y^{n-1} dy \quad [y := 2x-1] \\
&= \frac{n}{2} \left[\int_0^1 y^n dy + \int_0^1 y^{n-1} dy \right] \\
&= \frac{n}{2} \left(\frac{1}{n+1} + \frac{1}{n} \right) = 1 - \frac{1}{2(n+1)}.
\end{aligned} \tag{1}$$

Also,

$$\begin{aligned}
E\left(|\hat{\theta}/\theta|^2\right) &= 2n \int_{1/2}^1 x^2(2x-1)^{n-1} dx \\
&= n \int_0^1 \left(\frac{y+1}{2}\right)^2 y^{n-1} dy \quad [y := 2x-1] \\
&= \frac{n}{4} \int_0^1 (y^2 + 2y + 1) y^{n-1} dy \\
&= \frac{n}{4} \left[\int_0^1 y^{n+1} dy + 2 \int_0^1 y^n dy + \int_0^1 y^{n-1} dy \right] \\
&= \frac{n}{4} \left[\frac{1}{n+2} + \frac{2}{n+1} + \frac{1}{n} \right].
\end{aligned}$$

Note that $(n+1)^{-1} = n^{-1} - a_n$ where $a_n \approx n^{-2}$ as $n \rightarrow \infty$. Similarly, $(n+2)^{-1} = n^{-1} - b_n$ where $b_n \approx 2n^{-2}$ as $n \rightarrow \infty$. Therefore,

$$E\left(|\hat{\theta}/\theta|^2\right) = \frac{n}{4} \left[\frac{4}{n} - a_n - b_n \right] = 1 - c_n,$$

where $c_n = \frac{1}{4}\{na_n + nb_n\} \approx 3/(4n) \rightarrow 0$ as $n \rightarrow \infty$. This fact and (1) together yield

$$\begin{aligned}
\text{Var}(\hat{\theta}/\theta) &= 1 - c_n - \left[1 - \frac{1}{2(n+1)} \right]^2 \\
&= -c_n + \frac{1}{n+1} - \frac{1}{4(n+1)^2} \rightarrow 0,
\end{aligned}$$

as $n \rightarrow \infty$. Therefore, the Chebyshev inequality yields

$$\frac{\hat{\theta}}{\theta} - E\left[\frac{\hat{\theta}}{\theta}\right] \xrightarrow{P} 0 \quad \text{as } n \rightarrow \infty.$$

One more application of (1) shows that $\hat{\theta}/\theta \xrightarrow{P} 1$ as $n \rightarrow \infty$. This is equivalent to the desired consistency of $\hat{\theta}$.

- (c) Calculate the bias of the MLE. Use your computation to verify that the MLE is asymptotically unbiased [that is, show that $\text{bias}(\hat{\theta}) \rightarrow 0$ as $n \rightarrow \infty$].

Solution. We saw earlier in (1) that $E(\hat{\theta}/\theta) = 1 - \{2(n+1)\}^{-1}$. Therefore,

$$\text{Bias}(\hat{\theta}) = \theta - E(\hat{\theta}) = \frac{\theta}{2(n+1)}.$$

Clearly, this quantity converges to zero as $n \rightarrow \infty$.