Chapter 8

Martingales

8.18. Because $\mathcal{F}_{k+n-1} \subset \mathcal{F}_n$ for all $k, n \ge 1$, we may apply the tower property of conditional expectations to see that

$$\operatorname{E}[X_{k+n} | \mathcal{F}_n] = \operatorname{E}\left(\operatorname{E}[X_{k+n} | \mathcal{F}_{k+n-1}] \middle| \mathcal{F}_n\right) \ge \operatorname{E}[X_{k+n-1} | \mathcal{F}_n],$$

almost surely. Now apply induction to deduce the following: A.s.,

$$\operatorname{E}[X_{k+n} \,|\, \mathfrak{F}_n] \geqslant \operatorname{E}[X_{k+n-1} \,|\, \mathfrak{F}_n] \geqslant \cdots \geqslant \operatorname{E}[X_n \,|\, \mathfrak{F}_n] = X_n.$$

8.19. Suppose we could write our submartingale X as: $X_n = M_n + Z_n$ and $X_n = M'_n + Z'_n$, where M and M' are martingales, and Z and Z' are previsible increasing processes. Because M - M' defines a martingale, this proves that $Z_n - Z'_n$ defines a martingale also. That is,

$$E[Z_n - Z'_n | \mathcal{F}_{n-1}] = Z_{n-1} - Z'_{n-1}$$
 a.s

But Z - Z' is previsible. So the preceding also equals $Z_n - Z'_n$ a.s. This proves that for all $n \ge 1$, $Z_n - Z_{n-1} = Z'_n - Z'_{n-1}$ a.s. Thus, for all $m \ge 1$, $Z_m = Z_1 + \sum_{j=2}^m (Z_j - Z_{j-1}) = Z_1 + \sum_{j=2}^m (Z'_j - Z'_{j-1}) = Z_1 - Z'_1 + Z'_m$ a.s. Thus, if we insist that $Z_1 = Z'_1 = E[X_1]$, as was the case in the proof of Theorem 8.20, then $Z_m = Z'_m$ a.s. for all m.

- **8.20.** If X is bounded in L¹(P) then $E[X_n^+] \leq E[X_n^+] + E[X_n^-] = ||X_n||_1$ is bounded. For the converse note that $E[X_n^-] = E[X_n^+] - E[X_n] = E[X_n^+] - E[X_1]$ by the martingale property. This proves that $||X_n||_1 = 2E[X_n^+] - E[X_1]$ is bounded also.
- **8.26.** Without loss of generality, we assume that $Y \ge 0$; otherwise, we consider Y^+ and Y^- separately.

Let $\mathcal{F}_n := \sigma(X_n)$ and note that $\{\mathcal{F}_n\}_{n=1}^{\infty}$ is a filtration of σ -algebras in \mathcal{F} . [This is because of the nested structure of dyadic rationals.] By the

martingale convergence theorem, $Z := \lim_{n \to \infty} E(Y | X_n)$ exists a.s. and is finite a.s. Our goal is to prove that Z = E(Y | X).

Let I be a dyadic-rational interval, and note that if the length |I| is at least 2^{-N} , then $\{X \in I\} = \{X_n \in I\}$ for all $n \ge N$. Therefore,

$$\begin{split} \operatorname{E}\left[\operatorname{E}(Y | X_n); X \in \mathrm{I}\right] &= \operatorname{E}\left[\operatorname{E}(Y | X_n); X_n \in \mathrm{I}\right] \\ &= \operatorname{E}[Y; X_n \in \mathrm{I}] \\ &= \operatorname{E}[Y; X \in \mathrm{I}] \\ &= \operatorname{E}\left[\operatorname{E}(Y | X); X \in \mathrm{I}\right] \quad \text{for all } n \geq \mathrm{N}. \end{split}$$

Recall that $Z := \lim_{n\to\infty} E(Y|X_n)$ exists a.s. and is finite a.s. If we knew, additionally, that Y is a bounded random variable—say $|Y| \leq K$ a.s. for a constant K—then $|E(Y|X_n)| \leq E(|Y||X_n) \leq K$ by the conditional Jensen inequality. Consequently, the bounded convergence theorem would yield

$$\mathbf{E}\left[\mathsf{Z};\,\mathsf{X}\in\mathsf{I}\right]=\mathbf{E}\left[\mathbf{E}(\mathsf{Y}\,|\,\mathsf{X});\,\mathsf{X}\in\mathsf{I}\right],$$

for all dyadic-rational intervals I. Since dyadic rational intervals generate $\mathcal{B}(\mathbf{R})$, it follows that we could in fact choose any Borel set I in the preceding display, and therefore, it follows that E[Z; A] = E[E(Y|X); A] for all $A \in \sigma(X)$. Since Z and E(Y|X) are both $\sigma(X)$ -measurable, the uniqueness of conditional expectations yields Z = E(Y|X), as desired.

When Y is not bounded we can find a bounded random variable Y_K such that $||Y - Y_K||_1 \leq K^{-1}$ for all constants $K \geq 1$. The preceding shows that $\lim_{n\to\infty} E(Y_K | X_n) = E(Y_K | X)$ a.s. This convergence holds also in $L^1(P)$ because of the bounded convergence theorem.

But now we note from conditional Jensen inequality that

$$\|\mathrm{E}(\mathbf{Y} \,|\, \mathbf{X}_{n}) - \mathrm{E}(\mathbf{Y}_{K} \,|\, \mathbf{X}_{n})\|_{1} \leqslant \mathbf{K}^{-1},$$

and

$$\| E(Y|X) - E(Y_K|X) \|_1 \leq K^{-1}.$$

Therefore,

$$\left\| \operatorname{E}(Y \,|\, X_n) - \operatorname{E}(Y \,|\, X) \right\|_1 \leqslant 2K^{-1} + \left\| \operatorname{E}(Y_K \,|\, X_n) - \operatorname{E}(Y_K \,|\, X) \right\|_1 \to 2K^{-1},$$

as $n \to \infty$. Let $K \to \infty$ to conclude that $E(Y|X_n) \to E(Y|X)$ in $L^1(P)$. Since $E(Y|X_n) \to Z$ a.s., it follows that Z = E(Y|X) a.s.

Note that we have proved, for free, the additional fact that $E(Y|X_n) \rightarrow E(Y|X)$ in $L^1(P)$.

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