

8.26. Define \mathcal{F}_n to be the collection of all sets $X^{-1}(I)$, where I is a dyadic interval. Thus, $\{\mathcal{F}_n\}_{n=1}^{\infty}$ is a filtration, and $E[Y | X_n] = E[Y | \mathcal{F}_n]$. Because it is bounded in $L^1(\mathbb{P})$, $M_n := E[Y | X_n]$ has an a.s. limit M_{∞} . Evidently, M_{∞} is measurable with respect to $\mathcal{F}_{\infty} := \bigvee_{n=1}^{\infty} \mathcal{F}_n := \sigma(X)$.

First suppose in addition Y is bounded, so that $M_n \rightarrow M_{\infty}$ in $L^1(\mathbb{P})$ as well. Because conditional expectations are contractions on $L^1(\mathbb{P})$, this shows that $M_n = E[M_{n+k} | \mathcal{F}_n] \rightarrow E[M_{\infty} | \mathcal{F}_n]$ in $L^1(\mathbb{P})$, as $k \rightarrow \infty$. Thus, $E[M_{\infty} | \mathcal{F}_n] = M_n$ a.s. for all $n \geq 1$. Thus, for all $A \in \mathcal{F}_{\infty}$, $E[M_{\infty}; A] = E[M_n; A] = E[Y; A]$. This proves that $E[M_{\infty}; A] = E[Y; A]$ for all $A \in \mathcal{F}_{\infty}$. Therefore, $M_{\infty} = E[Y | \mathcal{F}_{\infty}] = E[Y | X]$.

Next suppose $Y \geq 0$. What we just proved shows that $E[Y \wedge k | X_n] \rightarrow E[Y \wedge k | X]$ a.s. as $n \rightarrow \infty$. By Doob's inequality, for all $\lambda > 0$,

$$\mathbb{P} \left\{ \sup_{n \geq 1} \left| E[Y | X_n] - E[Y \wedge k | X_n] \right| > \lambda \right\} \leq \frac{1}{\lambda} E[Y; Y > k].$$

In particular,

$$\mathbb{P} \left\{ \limsup_{n \rightarrow \infty} \left| E[Y | X_n] - E[Y | X] \right| > \lambda \right\} \leq \frac{1}{\lambda} E[Y; Y > k].$$

Let $k \rightarrow \infty$ to find that when Y is nonnegative, $\lim_{n \rightarrow \infty} E[Y | X_n] = E[Y | X]$. The general result follows from this and writing $Y = Y^+ - Y^-$.

8.27. Without loss of too much generality we consider the case where X is one-dimensional. Let X_n be as in **8.26**, and note that $E[Y | X_n]$ is a nonrandom Borel function of X_n . The result follows from **8.26**.