

Midterm Exam, Math 6010-1

Fall 2016

September 28, 2016

This is a 50-minute exam. You may use your textbook, as well as a calculator, but your work must be completely yours.

The exam is made of 3 questions in 4 pages, and is worth 40 points, total. Be sure to try all of the problems.

Partial credit is given only to carefully-written solutions.

1. (15 points in 3 equal parts) Which of the following are quadratic forms? Explain your reasoning, and identify the matrix of each quadratic form explicitly.

- (a) (5 points) $Q(\mathbf{x}) = \sum_{i=1}^n x_i$ for every $\mathbf{x} = (x_1, \dots, x_n)' \in \mathbb{R}^n$.

Solution. This is not a quadratic form. Here is why. Suppose to the contrary that there existed a matrix \mathbf{A} such that $Q(\mathbf{x}) = \mathbf{x}'\mathbf{A}\mathbf{x}$ for all $\mathbf{x} \in \mathbb{R}^n$. That is, suppose

$$Q(\mathbf{x}) = \sum_{i=1}^n \sum_{j=1}^n A_{i,j} x_i x_j.$$

Then, $(\partial^2 Q(\mathbf{x})) / (\partial x_i \partial x_j) = A_{i,j}$ for every $1 \leq i, j \leq n$. For us, however, $(\partial^2 Q(\mathbf{x})) / (\partial x_i \partial x_j) = 0$ for every i, j . This implies that $A_{i,j} = 0$ for all i, j , which means that $\mathbf{A} = \mathbf{0}$, an impossibility.

- (b) (5 points) $Q(\mathbf{x}) = \sum_{i=1}^n \sum_{j=1}^n (x_i - \bar{x})(x_j - \bar{x})$ for every $\mathbf{x} = (x_1, \dots, x_n)' \in \mathbb{R}^n$, where $\bar{x} := \frac{1}{n} \sum_{i=1}^n x_i$.

Solution. First, notice that $Q(\mathbf{x}) = \sum_{i=1}^n (x_i - \bar{x}) \cdot \sum_{j=1}^n (x_j - \bar{x}) = 0$. Thus, we see that Q is a quadratic form with trivial matrix $\mathbf{A} = \mathbf{0}$.

- (c) (5 points) $Q(\mathbf{x}) = \sum_{i=1}^{n-1} (x_{i+1} - x_i)^2$ for every $\mathbf{x} = (x_1, \dots, x_n)' \in \mathbb{R}^n$.

Solution. Write $Q = Q_1 + Q_2 - 2Q_3$, where

$$Q_1(\mathbf{x}) = \sum_{i=1}^{n-1} x_{i+1}^2, \quad Q_2(\mathbf{x}) = \sum_{i=1}^{n-1} x_i^2, \quad Q_3(\mathbf{x}) = \sum_{i=1}^{n-1} x_i x_{i+1}.$$

Each Q_j is a quadratic form, with respective matrices,

$$\mathbf{A}_1 = \begin{pmatrix} \mathbf{0}_{1 \times (n-1)} & 0 \\ \mathbf{I}_{(n-1) \times (n-1)} & \mathbf{0}_{(n-1) \times 1} \end{pmatrix}, \quad \mathbf{A}_2 = \begin{pmatrix} \mathbf{I}_{(n-1) \times (n-1)} & \mathbf{0}_{(n-1) \times 1} \\ \mathbf{0}_{1 \times (n-1)} & 0 \end{pmatrix},$$

and \mathbf{A}_3 is a matrix that is all ones but the super-diagonal entire; those super-diagonal entries are all ones; that is,

$$\mathbf{A}_3 = \begin{pmatrix} 0 & 1 & 0 & \cdots & 0 & 0 & 0 \\ 0 & 0 & 1 & \ddots & 0 & 0 & 0 \\ \vdots & \vdots & \ddots & \ddots & \vdots & \vdots & \vdots \\ 0 & 0 & 0 & \ddots & 1 & 0 & 0 \\ 0 & 0 & 0 & \cdots & 0 & 1 & 0 \\ 0 & 0 & 0 & \cdots & 0 & 0 & 1 \\ 0 & 0 & 0 & \cdots & 0 & 0 & 0 \end{pmatrix}.$$

Therefore, Q is a quadratic form with matrix $\mathbf{A}_1 - 2\mathbf{A}_3 + \mathbf{A}_2$.

2. (10 points total) The claim has been made that a certain random vector \mathbf{X} has a $N_3(\mathbf{0}, \mathbf{\Sigma})$ distribution, where

$$\mathbf{\Sigma} = \begin{pmatrix} 10 & 2 & 3 \\ 4 & 200 & 6 \\ 7 & 8 & 9 \end{pmatrix}.$$

- (a) (2 points) The eigenvalues of $\mathbf{\Sigma}$ are (approximately) 4.83, 13.87, and 200.3. Can the claim possibly be true?

Solution. It can potentially be true because $\mathbf{\Sigma}$ is positive semi-definite.

- (b) (5 points) Compute $E(\mathbf{X}'\mathbf{A}\mathbf{X})$ for the matrix $\mathbf{A} = \begin{pmatrix} 2 & 0 & 0 \\ 0 & -1 & 0 \\ 0 & 0 & 2 \end{pmatrix}$.

Solution. We know that

$$E(\mathbf{X}'\mathbf{A}\mathbf{X}) = \boldsymbol{\mu}'\mathbf{A}\boldsymbol{\mu} + \text{tr}(\mathbf{A}\mathbf{\Sigma}) = \sum_{i=1}^3 \sum_{j=1}^3 A_{i,j} \Sigma_{i,j} = \sum_{i=1}^3 A_{i,i} \Sigma_{i,i} = 20 - 200 + 18 = -162.$$

- (c) (3 points) Explain why your answer to (b) implies that \mathbf{A} is not positive semi-definite.

Solution. If \mathbf{A} were positive semidefinite, then $\mathbf{X}'\mathbf{A}\mathbf{X}$ would have to be ≥ 0 , and that would imply that $E(\mathbf{X}'\mathbf{A}\mathbf{X})$ would have to be ≥ 0 .

3. (15 points total) Let X_1, \dots, X_{10} be an independent random sample from a $N(0, 1)$ distribution, and define $\mathbf{Y} = (Y_1, Y_2)'$ as the vector of the following two “lagged averages”:

$$Y_1 = \frac{X_1 + \dots + X_5}{5} \quad \text{and} \quad Y_2 = \frac{X_2 + \dots + X_6}{5}.$$

- (a) (10 points) Compute $E(\mathbf{Y})$ and $\text{Var}(\mathbf{Y})$.

Solution. $E(Y_1) = E(Y_2) = 0$, therefore, $E(\mathbf{Y}) = \mathbf{0}$. Also, $\text{Var}(Y_1) = \text{Var}(Y_2) = 1/5$, and $\text{Cov}(Y_1, Y_2) = \frac{1}{36} \text{Var}(X_2 + \dots + X_5) = 4/36 = 1/9$. Therefore,

$$\text{Var}(\mathbf{Y}) = \begin{bmatrix} 1/5 & 1/9 \\ 1/9 & 1/5 \end{bmatrix}.$$

- (b) (5 points) Explain carefully why \mathbf{Y} has a bivariate normal distribution. Does \mathbf{Y} have a probability density?

Solution. \mathbf{Y} is a bivariate normal because $\mathbf{Z} = (X_1, \dots, X_6)'$ is a vector of i.i.d. standard normals, and

$$\mathbf{Y} = \begin{pmatrix} 1/5 & 1/5 & 1/5 & 1/5 & 1/5 & 0 \\ 0 & 1/5 & 1/5 & 1/5 & 1/5 & 1/5 \end{pmatrix} \mathbf{Z}.$$

The determinant of $\text{Var}(\mathbf{Y})$ is positive, so $\text{Var}(\mathbf{Y})$ is positive definite and hence \mathbf{Y} has a pdf. Indeed,

$$\det[\text{Var}(\mathbf{Y})] = (1/25) - (1/81) = 39/(25 \times 81) > 0.$$