

Math 6010

Solutions to homework 2

3, p. 15. Recall that $\sum_{i=1}^n (x_i - \bar{x})^2 = \mathbf{x}' \mathbf{A} \mathbf{x}$ for every \mathbf{x} , where

$$\mathbf{A} := \begin{pmatrix} 1 - \frac{1}{n} & -\frac{1}{n} & \cdots & -\frac{1}{n} & -\frac{1}{n} \\ \vdots & \vdots & \ddots & \vdots & \vdots \\ -\frac{1}{n} & -\frac{1}{n} & \cdots & -\frac{1}{n} & 1 - \frac{1}{n} \end{pmatrix}.$$

Therefore, $E(Q_1) = \text{tr}(\mathbf{A}\mathbf{\Sigma}) + \boldsymbol{\mu}' \mathbf{A} \boldsymbol{\mu}$. Because $\boldsymbol{\mu}' \mathbf{A} \boldsymbol{\mu} = \sum_{i=1}^n (\mu_i - \bar{\mu})^2$ and the μ_i 's are all the same it follows that $\boldsymbol{\mu}' \mathbf{A} \boldsymbol{\mu} = 0$ and therefore $E(Q_1) = \text{tr}(\mathbf{A}\mathbf{\Sigma}) = \sum_i \sum_j A_{i,j} \Sigma_{i,j}$. Because $\Sigma_{i,j} = 0$ when $j > i + 1$, and since $\mathbf{\Sigma}$ and \mathbf{A} are both symmetric,

$$\begin{aligned} E(Q_1) &= \sum_i \sum_j A_{i,j} \Sigma_{i,j} = \sum_{i=1}^n A_{i,i} \Sigma_{i,i} + 2 \sum_{i=1}^n A_{i,i+1} \Sigma_{i,i+1} \\ &= \left(1 - \frac{1}{n}\right) \sum_{i=1}^n \Sigma_{i,i} - \frac{2}{n} \sum_{i=1}^n \Sigma_{i,i+1}. \end{aligned}$$

This does the job for Q_1 . As regards Q_2 , we appeal to Example 1.8 (p. 10) to see that

$$\begin{aligned} Q_2 &= 2 \sum_{i=1}^n X_i^2 - X_1^2 - X_n^2 - 2 \sum_{i=1}^{n-1} X_i X_{i+1} + (X_1 - X_n)^2 \\ &= 2 \sum_{i=1}^n X_i^2 - 2 \sum_{i=1}^{n-1} X_i X_{i+1} - 2X_1 X_n. \end{aligned}$$

Because $E(X_1 X_n) = E(X_1)E(X_n) = \mu^2$ [I am assuming that $n > 3$; otherwise the result is silly], it follows that

$$E(Q_2) = 2 \sum_{i=1}^n E(X_i^2) - 2 \sum_{i=1}^{n-1} E(X_i X_{i+1}) - 2\mu^2.$$

Now, $E(X_i^2) = \text{Var}(X_i) + (E X_i)^2 = \Sigma_{i,i} + \mu^2$, and $E(X_i X_{i+1}) = \Sigma_{i,i+1} + \mu^2$ similarly. Therefore,

$$E(Q_2) = 2 \sum_{i=1}^n \Sigma_{i,i} + 2n\mu^2 - 2 \sum_{i=1}^{n-1} \Sigma_{i,i+1} - 2(n-1)\mu^2 - 2\mu^2 = 2 \sum_{i=1}^n \Sigma_{i,i} - 2 \sum_{i=1}^{n-1} \Sigma_{i,i+1}.$$

Therefore,

$$\begin{aligned}
\mathbb{E}(3Q_1 - Q_2) &= 3 \left[\left(1 - \frac{1}{n}\right) \sum_{i=1}^n \Sigma_{i,i} - \frac{2}{n} \sum_{i=1}^n \Sigma_{i,i+1} \right] - \left[2 \sum_{i=1}^n \Sigma_{i,i} - 2 \sum_{i=1}^{n-1} \Sigma_{i,i+1} \right] \\
&= \left(1 - \frac{3}{n}\right) \sum_{i=1}^n \Sigma_{i,i} + \left(2 - \frac{6}{n}\right) \sum_{i=1}^{n-1} \Sigma_{i,i+1} \\
&= \frac{n-3}{n} \sum_{i=1}^n \Sigma_{i,i} + 2 \left(\frac{n-3}{n}\right) \sum_{i=1}^{n-1} \Sigma_{i,i+1}.
\end{aligned}$$

Divide both sides by $n(n-3)$ in order to see that

$$\mathbb{E} \left(\frac{3Q_1 - Q_2}{n(n-3)} \right) = \frac{1}{n^2} \sum_{i=1}^n \Sigma_{i,i} + \frac{2}{n^2} \sum_{i=1}^{n-1} \Sigma_{i,i+1}.$$

Finally, recall that

$$\text{Var}(X_1 + \cdots + X_n) = \sum_{i=1}^n \sum_{j=1}^n \text{Cov}(X_i, X_j) = \sum_{i=1}^n \Sigma_{i,i} + 2 \sum_{i=1}^{n-1} \Sigma_{i,i+1}.$$

[The sum is the sum of the diagonal entries plus twice the sum of the entires that lie above the diagonal, by symmetry.] Therefore,

$$\text{Var}(\bar{X}) = \frac{1}{n^2} \text{Var}(X_1 + \cdots + X_n) = \frac{1}{n^2} \sum_{i=1}^n \Sigma_{i,i} + \frac{2}{n^2} \sum_{i=1}^{n-1} \Sigma_{i,i+1},$$

which as we have seen is equal to the expectation of $(3Q_1 - Q_2)/n(n-3)$.

4, p. 16. The matrix of the quadratic form $(x_1 - x_2)^2 + (x_2 - x_3)^2 + (x_3 - x_1)^2$ is

$$\mathbf{A} := \begin{pmatrix} 2 & -1 & -1 \\ -1 & 2 & -1 \\ -1 & -1 & 2 \end{pmatrix}.$$

According to Theorem 1.6 (p. 10), the variance of $(X_1 - X_2)^2 + (X_2 - X_3)^2 + (X_1 - X_3)^2 = \mathbf{X}' \mathbf{A} \mathbf{X}$ is

$$(\mu_4 - 3\mu_2^2) \mathbf{a}' \mathbf{a} + 2\mu_2^2 \text{tr}(\mathbf{A}^2),$$

because the means [i.e., the θ 's in that theorem] are all zero. Here, $\mathbf{a} := (2, 2, 2)'$ is the diagonals-vector of \mathbf{A} ,

$$\mu_2 = \mathbb{E}(X_i^2) = \int_{-1}^1 \frac{1}{2} x^2 dx = \frac{1}{3},$$

and

$$\mu_4 = E(X_i^4) = \int_{-1}^1 \frac{1}{2} x^4 dx = \frac{1}{5}.$$

Therefore, $(\mu_4 - 3\mu_2^2)\mathbf{a}'\mathbf{a} = -\frac{24}{15} = -\frac{8}{5}$. Since \mathbf{A} is symmetric,

$$\text{tr}(\mathbf{A}^2) = \sum_{i=1}^3 \sum_{j=1}^3 A_{i,j} A_{i,j} = \sum_{i=1}^3 \sum_{j=1}^3 A_{i,j}^2 = 18.$$

Therefore, the answer to this question is $-\frac{8}{5} + 4 = \frac{12}{5}$.

5, p. 16. Observe that $\mathbf{X}'\mathbf{A}\mathbf{X}$ and $\mathbf{X}'\mathbf{B}\mathbf{X}$ are 1-dimensional random variables. Consequently,

$$\begin{aligned} \mathbf{X}'\mathbf{A}\mathbf{X} \cdot \mathbf{X}'\mathbf{B}\mathbf{X} &= \sum_{i=1}^n \sum_{j=1}^n X_i A_{i,j} X_j \cdot \sum_{k=1}^n \sum_{l=1}^n X_k B_{k,l} X_l \\ &= \sum_{i=1}^n \sum_{j=1}^n \sum_{k=1}^n \sum_{l=1}^n A_{i,j} B_{k,l} X_i X_j X_k X_l. \end{aligned}$$

Take expectations to see that

$$\begin{aligned} E[\mathbf{X}'\mathbf{A}\mathbf{X}, \mathbf{X}'\mathbf{B}\mathbf{X}] &= E(\mathbf{X}'\mathbf{A}\mathbf{X} \cdot \mathbf{X}'\mathbf{B}\mathbf{X}) \\ &= \sum_{i=1}^n \sum_{j=1}^n \sum_{k=1}^n \sum_{l=1}^n A_{i,j} B_{k,l} E(X_i X_j X_k X_l). \end{aligned}$$

Because the X_i 's are independent and have mean and third moment zero, all of these terms are zero except in the following cases: (i) When $i = j = k = l$; (ii) $i = j \neq k = l$; (iii) $i = k \neq j = l$; (iv) $i = l \neq j = k$. In case (i), $E(X_i X_j X_k X_l) = E(X_1^4) = 3\sigma^4$. In the remaining cases (i)–(iv), $E(X_i X_j X_k X_l) = |E(X_1^2)|^2 = \sigma^4$. Therefore,

$$\begin{aligned} E[\mathbf{X}'\mathbf{A}\mathbf{X}, \mathbf{X}'\mathbf{B}\mathbf{X}] &= 3\sigma^4 \sum_{i=1}^n A_{i,i} B_{i,i} + \sigma^4 \sum_{i=1}^n \sum_{\substack{k=1 \\ k \neq i}}^n A_{i,i} B_{k,k} \\ &\quad + \sigma^4 \sum_{i=1}^n \sum_{\substack{j=1 \\ j \neq i}}^n A_{i,j} B_{i,j} + \sigma^4 \sum_{i=1}^n \sum_{\substack{j=1 \\ j \neq i}}^n A_{i,j} B_{j,i}, \end{aligned}$$

considering each case separately in order from (i)–(iv). Because \mathbf{A} and \mathbf{B} are symmetric, the last two quantities are equal. This reduces the computation to the following:

$$\begin{aligned} E[\mathbf{X}'\mathbf{A}\mathbf{X}, \mathbf{X}'\mathbf{B}\mathbf{X}] &= 3\sigma^4 \sum_{i=1}^n A_{i,i} B_{i,i} + \sigma^4 \sum_{i=1}^n \sum_{\substack{k=1 \\ k \neq i}}^n A_{i,i} B_{k,k} \\ &\quad + 2\sigma^4 \sum_{i=1}^n \sum_{\substack{j=1 \\ j \neq i}}^n A_{i,j} B_{i,j}. \end{aligned}$$

Now,

$$\begin{aligned}\sum_{i=1}^n \sum_{\substack{j=1 \\ j \neq i}}^n A_{i,j} B_{i,j} &= \sum_{i=1}^n \sum_{j=1}^n A_{i,j} B_{i,j} - \sum_{i=1}^n A_{i,i} B_{i,i} \\ &= \text{tr}(\mathbf{AB}) - \sum_{i=1}^n A_{i,i} B_{i,i}.\end{aligned}$$

Consequently,

$$\begin{aligned}\text{E}[\mathbf{X}'\mathbf{AX}, \mathbf{X}'\mathbf{BX}] &= \sigma^4 \sum_{i=1}^n A_{i,i} B_{i,i} + \sigma^4 \sum_{i=1}^n \sum_{\substack{k=1 \\ k \neq i}}^n A_{i,i} B_{k,k} \\ &\quad + 2\sigma^4 \text{tr}(\mathbf{AB}) \\ &= \sigma^4 \sum_{i=1}^n \sum_{k=1}^n A_{i,i} B_{k,k} + 2\sigma^4 \text{tr}(\mathbf{AB}) \\ &= \sigma^4 \text{tr}(\mathbf{A}) \text{tr}(\mathbf{B}) + 2\sigma^4 \text{tr}(\mathbf{AB}) \\ &= \text{E}(\mathbf{X}'\mathbf{AX}) \text{E}(\mathbf{X}'\mathbf{BX}) + 2\sigma^4 \text{tr}(\mathbf{AB}),\end{aligned}$$

thanks to Theorem 1.5. This shows readily that

$$\text{Cov}[\mathbf{X}'\mathbf{AX}, \mathbf{X}'\mathbf{BX}] = 2\sigma^4 \text{tr}(\mathbf{AB}).$$