

**Math 5090–001, Fall 2009**  
**Solutions to Assignment 2**

**Chapter 11, Problem 17.** (a) The cdf of  $X \sim \text{GEO}(p)$  is

$$G(s; p) = \sum_{j=1}^s p(1-p)^{j-1} = p \sum_{i=0}^{s-1} (1-p)^i = p \left[ \frac{1 - (1-p)^s}{1 - (1-p)} \right] = 1 - (1-p)^s,$$

for  $s = 1, 2, \dots$ . Viewed as a function of  $p$ ,  $G(s; p)$  is increasing for every integer  $s = 1, 2, \dots$ . In fact,

$$\frac{\partial}{\partial p} G(s; p) = s(1-p)^{s-1} \geq 0.$$

Therefore, the functions  $h_1$  and  $h_2$  of Theorem 11.4.3 are decreasing. A conservative one-sided  $100(1 - \alpha)\%$  lower CI for  $p$  is  $(\theta_L, 1]$ , where we find  $\theta_L$  by solving  $G(X; \theta_L) = \alpha$ . That is, find  $\theta_L$  such that

$$1 - (1 - \theta_L)^X = \alpha \quad \Leftrightarrow \quad \theta_L = 1 - (1 - \alpha)^{1/X}.$$

(b) The answer becomes  $\theta_L = 1 - 0.9^{1/5} \approx 0.02085$ . So the confidence interval is  $\approx (0.029, 1]$ .

(c) The sum of  $n$  independent  $\text{GEO}(p)$ 's is negative binomial; i.e.,

$$G(s; p) = \sum_{k=n}^s \binom{k-1}{n-1} p^n (1-p)^{k-n},$$

for  $s \geq n$ , an integer. If  $s$  is fixed, then

$$\frac{\partial}{\partial p} G(s; p) = \sum_{k=n}^s \binom{k-1}{n-1} \left\{ np^{n-1}(1-p)^{k-n} + (k-n)p^n(1-p)^{k-n-1} \right\} \geq 0.$$

So, once again,  $G(s, p)$  increases with  $p$ , for every fixed integer  $s \geq n$ . And  $h_1$  and  $h_2$  are decreasing; so the CI is  $(\theta_L, 1]$ , where  $\theta_L$  is the solution to  $G(S; \theta_L) = \alpha$ , where  $S := X_1 + \dots + X_n$  is the sum of the  $n$  i.i.d. geometrics.

**Chapter 11, Problem 27. (a)** [*I graded only this part.*] The length of the CI is

$$\{\Phi^{-1}(\alpha_1) - \Phi^{-1}(1 - \alpha_2)\} \cdot \frac{\sigma}{\sqrt{n}}$$

Our goal is to minimize the preceding function of  $\alpha_1, \alpha_2 \geq 0$ , subject to  $\alpha_1 + \alpha_2 = \alpha$ . Equivalently, we must minimize

$$h(\alpha_1, \alpha - \alpha_1) := \Phi^{-1}(\alpha_1) - \Phi^{-1}(1 - \alpha + \alpha_1) \quad \text{over all } \alpha_1 \in (0, 1).$$

Now, elementary calculus tells us that

$$\begin{aligned} \frac{\partial}{\partial \alpha_1} h(\alpha_1, \alpha - \alpha_1) &= \frac{\partial}{\partial \alpha_1} \Phi^{-1}(\alpha_1) - \frac{\partial}{\partial \alpha_1} \Phi^{-1}(1 - \alpha + \alpha_1) \\ &= \frac{1}{\phi(\Phi^{-1}(\alpha_1))} - \frac{1}{\phi(\Phi^{-1}(1 - \alpha + \alpha_1))}, \end{aligned}$$

where  $\phi$  is the  $N(0, 1)$  density function. Set the preceding to zero to find that the optimal choice of  $\alpha_1$  solves

$$\phi(\Phi^{-1}(\alpha_1)) = \phi(\Phi^{-1}(1 - \alpha + \alpha_1)).$$

Inspect the curve  $\phi$  to see that if  $\phi(a) = \phi(b)$  and  $a \neq b$ , then  $a = -b$ , necessarily! This means that

$$\Phi^{-1}(\alpha_1) = -\Phi^{-1}(1 - \alpha + \alpha_1).$$

In other words [check the graph of  $\Phi$ ],  $\alpha_1 = \alpha - \alpha_1$ , and hence  $\alpha_1 = \alpha/2$ . In order to complete this computation, you need to show that the second derivative of  $h(\alpha_1, \alpha - \alpha_1)$  [with respect to  $\alpha_1$ ] is positive at  $\alpha_1 = \alpha/2$ . I will let you do this yourself.

**(b)** Yes; replace  $\sigma$  by  $S$  and  $\Phi^{-1}(\alpha_j)$  by  $F^{-1}(\alpha_j)$ , where  $F$  denotes the cdf of a  $t$  distribution with  $n - 1$  df. In this way, we find that the optimal  $\alpha_1$  solves

$$F'(F^{-1}(\alpha_1)) = F'(F^{-1}(1 - \alpha + \alpha_1)).$$

Since the shape of  $F$  is similar to  $\Phi$ , it has the same properties, and so  $\alpha_1 = \alpha/2$  follows from the same considerations as before [you should check the details!].

(c) Because  $E(S^2) = \sigma^2$ , the expected length of the general form of the CI (11.3.6) is

$$2(n-1)\sigma^2 \left[ \frac{1}{q_1} - \frac{1}{q_2} \right],$$

where  $F(q_1) = \alpha_1$  and

$$F(q_2) = \alpha_2 = 1 - \alpha + \alpha_1 = \text{constant} + F(q_1), \quad (\star)$$

and  $F$  is the cdf of a  $\chi^2(n-1)$ . Let  $f := F'$ , as usual. We can differentiate  $(\star)$   $[d/dq_1]$  to find that

$$f(q_2) \frac{dq_2}{dq_1} = f(q_1) \quad \Rightarrow \quad \frac{dq_2}{dq_1} = \frac{f(q_1)}{f(q_2)}.$$

But the expected length is proportional to

$$h(q_1, q_2) = \frac{1}{q_1} - \frac{1}{q_2};$$

in fact, we saw that the constant of proportionality is  $2(n-1)\sigma^2$ . Want to set  $\partial h / \partial q_1 = 0$ ;

$$\frac{\partial h}{\partial q_1} = -\frac{1}{q_1^2} + \frac{1}{q_2^2} \frac{dq_2}{dq_1} = -\frac{1}{q_1^2} + \frac{1}{q_2^2} \frac{f(q_1)}{f(q_2)}.$$

Set this equal to zero to find that the optimal choice must solve

$$\frac{q_2^2}{q_1^2} = \frac{f(q_1)}{f(q_2)}.$$

If the equal-tailed answer did this for all  $\alpha \in (0, 1/2)$ , then  $f$  would have to be symmetric; and it is not.

**Chapter 11, Problem 30.** (a) Since  $\bar{X}_n$  is asymptotically  $N(\mu, \mu/n)$  and  $g(\mu) = \sqrt{\mu}$ , it follows that  $c = \sqrt{\mu}$  and  $g'(\mu) = 1/(2\sqrt{\mu})$ . Therefore,  $|cg'(\mu)|^2$  is a constant

[namely, 1/4] and so  $g$  is a variance-stabilizing transformation. And by Theorem 7.7.6 (p. 249),

$$\sqrt{\bar{X}_n} \stackrel{d}{\approx} N\left(\sqrt{\mu}, \frac{1}{4n}\right).$$

(b) We know that for  $n$  large,

$$P\left\{\left|\sqrt{\bar{X}_n} - \sqrt{\mu}\right| \leq \frac{z_{1-(\alpha/2)}}{2\sqrt{n}}\right\} \approx 1 - \alpha.$$

Therefore, an approx.  $100(1 - \alpha)\%$  CI for  $\sqrt{\mu}$  is

$$\left(\sqrt{\bar{X}_n} - \frac{z_{1-(\alpha/2)}}{2\sqrt{n}}, \sqrt{\bar{X}_n} + \frac{z_{1-(\alpha/2)}}{2\sqrt{n}}\right),$$

and for  $\mu$ , it translates to

$$\left(\left|\sqrt{\bar{X}_n} - \frac{z_{1-(\alpha/2)}}{2\sqrt{n}}\right|^2, \left|\sqrt{\bar{X}_n} + \frac{z_{1-(\alpha/2)}}{2\sqrt{n}}\right|^2\right).$$

(c) By the central limit theorem,

$$\frac{\sqrt{n} [\bar{X}_n - \theta]}{\theta} \stackrel{d}{\rightarrow} N(0, 1).$$

So,  $c(\theta) = \theta$ , and we seek to find a function  $g(\theta)$  such that  $|c(\theta)g'(\theta)|^2 = \theta^2|g'(\theta)|^2$  is a constant. Set  $g(\theta) := \ln \theta$  to see that indeed  $\theta^2|g'(\theta)|^2 = 1$ . Theorem 7.7.6 (p. 249) then tells us that

$$\sqrt{n} [\ln \bar{X}_n - \ln \theta] \stackrel{d}{\rightarrow} N(0, 1).$$

So an approximate  $100(1 - \alpha)\%$  CI for  $\ln \theta$  is

$$\left(\ln \bar{X}_n - \frac{z_{1-(\alpha/2)}}{\sqrt{n}}, \ln \bar{X}_n + \frac{z_{1-(\alpha/2)}}{\sqrt{n}}\right).$$

And an approximate  $100(1 - \alpha)\%$  CI for  $\theta$  is

$$(\bar{X}_n e^{-Q}, \bar{X}_n e^Q).$$

where  $Q := z_{1-(\alpha/2)}/\sqrt{n}$ .