

Solutions to Midterm 4

Math 5010-1, Spring 2007, University of Utah

1. Suppose X is a standard normal random variable. What is the density of $1/X^2$?

Solution: Let $Y = 1/X^2$ and note that for all $a > 0$,

$$\begin{aligned} F_Y(a) &= P\{X^{-2} \leq a\} = P\{X^2 > a^{-1}\} \\ &= 1 - P\{X^2 \leq a^{-1}\} = 1 - P\{|X| \leq 1/\sqrt{a}\} \\ &= 1 - (F_X(1/\sqrt{a}) - F_X(-1/\sqrt{a})) \\ &= 1 - (2F_X(1/\sqrt{a}) - 1) = 2 - 2F_X(1/\sqrt{a}). \end{aligned}$$

Therefore,

$$\begin{aligned} f_Y(a) &= -2f_X(1/\sqrt{a}) \frac{d}{da} \left(\frac{1}{\sqrt{a}} \right) \\ &= -2f_X(1/\sqrt{a}) \times \left(-\frac{1}{2a^{3/2}} \right) \\ &= \frac{1}{a^{3/2}} f_X(1/\sqrt{a}) = \frac{1}{a^{3/2}\sqrt{2\pi}} \exp\left(-\frac{1}{2a}\right). \end{aligned}$$

If $a \leq 0$, then $f_Y(a) = 0$.

2. Compute the moment generating function of X , when X is a random variable with each of the following densities:

(a) $f(x) = \frac{1}{2}e^{-|x|}$.

Solution: For all t ,

$$\begin{aligned} M(t) &= \int_{-\infty}^{\infty} e^{tx} \frac{e^{-|x|}}{2} dx \\ &= \int_0^{\infty} \frac{e^{tx-x}}{2} dx + \int_{-\infty}^0 \frac{e^{tx+x}}{2} dx \\ &= \int_0^{\infty} \frac{e^{tx-x}}{2} dx + \int_0^{\infty} \frac{e^{-tx-x}}{2} dx \\ &= \int_0^{\infty} \frac{e^{-x(1-t)}}{2} dx + \int_0^{\infty} \frac{e^{-x(1+t)}}{2} dx. \end{aligned}$$

If $t \geq 1$, then the first integral blows up. If $t \leq -1$, then the second integral blows up. Therefore, we have a finite $M(t)$ if and only if $|t| < 1$. In that case,

$$M(t) = \frac{1}{2(1-t)} + \frac{1}{2(1+t)} = \frac{1}{1-t^2}.$$

(b) $f(x) = 1 - |x|$ if $-1 < x < 1$; else, $f(x) = 0$.

Solution: For all real numbers t ,

$$\begin{aligned} M(t) &= \int_{-1}^1 (1 - |x|) e^{tx} dx = \int_{-1}^1 e^{tx} dx - \int_{-1}^1 |x| e^{tx} dx \\ &= \frac{e^t - e^{-t}}{t} - \int_0^1 x e^{tx} dx - \int_{-1}^0 (-x) e^{tx} dx \\ &= \frac{e^t - e^{-t}}{t} - \int_0^1 x e^{tx} dx - \int_0^1 x e^{-tx} dx. \end{aligned}$$

Integrate by parts to find that for all real numbers s ,

$$\begin{aligned} \int_0^1 x e^{sx} dx &= \int_0^1 \underbrace{x}_u \underbrace{e^{sx}}_{v'} dx = uv \Big|_0^1 - \int_0^1 \underbrace{\frac{1}{s} e^{sx}}_v \underbrace{1}_{u'} dx \\ &= \frac{e^s}{s} - \frac{1}{s} \int_0^1 e^{sx} dx = \frac{e^s}{s} - \frac{1 - e^s}{s^2}. \end{aligned}$$

Use this, once with $s = t$ and once with $s = -t$ to find an answer.

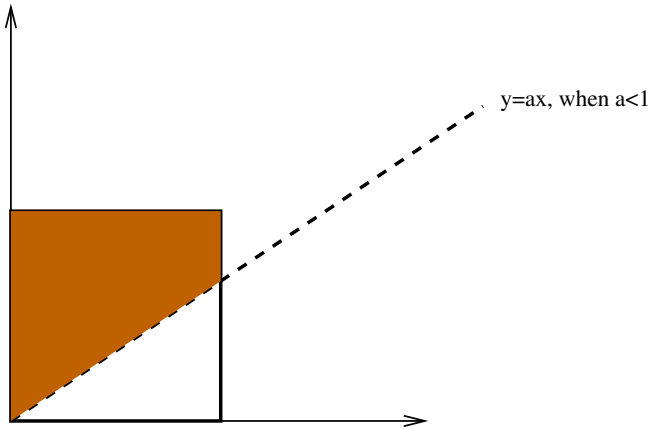


Figure 1: The case where $a < 1$

3. Suppose (X, Y) is distributed uniformly on the square with corners at $(0, 0)$, $(1, 0)$, $(0, 1)$, and $(1, 1)$. Compute $P\{Y > aX\}$ for all real numbers a .

Solution: By definition, the joint density function f of (X, Y) is

$$f(x, y) = \begin{cases} 1 & \text{if } 0 < x < 1 \text{ and } 0 < y < 1, \\ 0 & \text{otherwise.} \end{cases}$$

Therefore, we see to find $\iint_{y > ax} f(x, t) dx dy$. The computation depends on the value of a .

If $a = 1$, then $P\{Y > X\} = P\{Y > aX\} = 1/2$ by symmetry. Therefore, there are two cases to consider (Figures 1 and 2):

If $a < 1$, then according to Figure 1,

$$P\{Y > aX\} = \int_0^1 \int_{ax}^1 dy dx = \int_0^1 (1 - ax) dx = 1 - \frac{a}{2}.$$

If, on the other hand, $a > 1$, then according to Figure 2,

$$P\{Y > aX\} = \int_0^1 \int_0^{y/a} dx dy = \frac{1}{a} \int_0^1 y dy = \frac{1}{2a}.$$

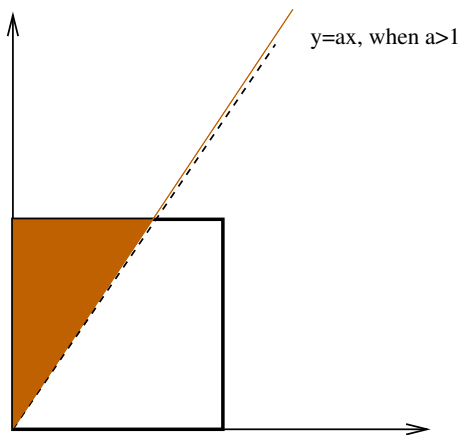


Figure 2: The case where $a > 1$

4. Find EX , when (X, Y) is jointly distributed with density

$$f(x, y) = \begin{cases} 8xy & \text{if } 0 < y < x < 1, \\ 0 & \text{otherwise.} \end{cases}$$

You must do this by first computing f_X .

Solution: Integrate out the y variable first: If $0 < x < 1$, then

$$f_X(x) = \int_0^x 8xy \, dy = 4x^3.$$

Else, $f_X(x) = 0$. Consequently,

$$EX = \int_0^1 4x^4 \, dx = \frac{4}{5}.$$