1. In 2005, the United Nations reported that approximately 50.74 percent of the adults in the US are women. Suppose we take a random sample of 100 people, without replacement, from this population. What are the odds that there are between 50 and 51 women in this sample?

Solution: Let N denote the population of the United States. Define W to be the total number of women, so that $W \approx 0.5074N$. Then the probability that there are 50 or 51 women is

$$\frac{\binom{W}{50}\binom{N-W}{50}}{\binom{N}{100}} + \frac{\binom{W}{51}\binom{N-W}{49}}{\binom{N}{100}}.$$

There is a good approximation to this. Since N is very large compared to the sample size of 100, sampling with replacement is pretty much the same as sampling without replacement. Therefore, the answer is close to the binomial probabilities,

$$\binom{100}{50} 0.5074^{50} \times 0.4926^{50} + \binom{100}{51} 0.5074^{51} \times 0.4926^{49} \approx 0.158.$$

There is a sense that this can be proved, but that is another matter.

2. There are k coins on the table. Coin i tosses heads with probability i/k for every i = 1,..., k. You choose one at random—all are equally likely—and toss it n times independently. It turns up heads. What is the probability that you had chosen coin k?

Solution: Let C_i denote the event that the *i*th coin is selected. Let H_n denote the event that the first tosses are all heads. By independence, $P(H_n | C_i) = (i/k)^n$. Therefore, we can apply the Bayes theorem to find that

$$P(C_k \mid H_n) = \frac{P(H_n \mid C_k)P(C_k)}{\sum_{i=1}^k P(H_n \mid C_i)P(C_i)} = \frac{\frac{1}{k}}{\sum_{i=1}^k \left(\frac{i}{k}\right)^n \left(\frac{1}{k}\right)} = \frac{1}{\sum_{i=1}^k \left(\frac{i}{k}\right)^n}$$

If you do this for n = 1, then that is acceptable. In that case, things simplify because $\sum_{i=1}^{k} (i/k) = \frac{1}{2}(k+1)$. Therefore, $P(C_k \mid H_1) = 2/(k+1)$ in that case.

3. There are n light bulbs in a storage facility, k of which are not functional. You select r at random, and without replacement. What is the probability that you select ℓ non-functional bulb? You may assume that all light bulbs are equally likely to be selected.

Solution: The probability is

$$\frac{\binom{k}{l}\binom{n-k}{r-l}}{\binom{n}{r}},$$

assuming, of course, that k, l, n, r are integers that satisfy: (i) $0 \le l \le k$; (ii) $0 \le r - l \le n - k$; and (iii) $0 \le r \le n$. Otherwise, the probability is zero.

4. A fair die is cast 112 times. What is the probability that we roll two dots at least twice?

Solution: I will state things in the language of binomial random variables because it will hopefully make things more clear. [Words to the wise: Binomials are not really needed here, though.]

Call it a success every time you roll two dots. The question asks to find $P\{X \ge 2\}$, where X is the number of successes. Clearly, X = Bin(112, 1/6). Therefore,

$$P\{X \ge 2\} = \sum_{j=2}^{112} {\binom{112}{j}} \left(\frac{1}{6}\right)^j \left(\frac{5}{6}\right)^{112-j}$$
$$= 1 - \left[{\binom{112}{0}} \left(\frac{1}{6}\right)^0 \left(\frac{5}{6}\right)^{112-0} + {\binom{112}{1}} \left(\frac{1}{6}\right)^1 \left(\frac{5}{6}\right)^{112-1} \right].$$

This can be simplified. Please find better things to do with your time though.

- 5. A fair coin is tossed, independently, countably many times.
 - (a) Carefully write down a sample space Ω for this experiment.

Solution: All possible infinite sequences of heads and tails.

(b) Prove, carefully using the laws of probability, that every element of Ω is possible, but has probability zero.

Solution: "Possible" is easy, since you can write any sequence in Ω down inductively. Let S_n denote the first n steps of the sequence. By independence, $P(S_n) = (1/2)^n$. Note that $S_n \subseteq S_{n-1}$ for all $n \ge 2$. Therefore, $0 = \lim_{n \to \infty} P(S_n)$ is the same as the probability of $\bigcap_{n=1}^{\infty} S_n$, which is the event that the entire sequence appears.